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Genetics of Parenting: The Power of the Dark Side

Developmental Psychology (brief report)

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Abstract (219 words)

Reviews of behavioral genetic studies note that ‘control’ aspects of parenting yield low estimates of heritability, while ‘affective’ aspects (parental feelings) yield moderate estimates.  Research to date has not specifically considered whether positive and negative aspects of parenting – for both feelings and control – may explain these etiological distinctions.  We addressed this issue using parent reports of parenting in a large UK twin sample, at ages 9 (N=2260 twin pairs), 12 (N=3850 twin pairs) and 14 years (N=2293 twin pairs).  Our findings supported previous work indicating that parental feelings show greater heritability (h2) than control (across all ages, mean h2 feelings=.42, control=.13). Of specific interest is our novel finding that for control as well as for feelings, the heritability for negative aspects of parenting was greater than for positive aspects (e.g., across all ages, mean h2 total negativity=.44; total positivity=.12).  Results across the three ages using common pathway models for all scales further endorsed our hypotheses. Previous research has shown that children’s genetically driven characteristics elicit parenting; our pattern of our results suggests that what is critical is the ‘dark’ side of these characteristics for eliciting negativity from parents, whether feelings towards the child or control strategies are considered. Improving our understanding of how the environment is shaped by the dark side is important theoretically, and, ultimately, for targeting intervention.

KEYWORDS: behavior genetics; parental control; parental feelings; parental negativity; parental positivity

Genetics of Parenting: The Power of the Dark Side

As first described by Bell (1968), parenting is an interactional process in which child characteristics influence parenting behaviors as well as the other way around. In a then controversial presidential address to the Society for Research in Child Development in 1991 (Scarr, 1992), behavioral genetic evidence was used to make the case that children’s experiences are dependent on their genetic propensities. Three types of so-called genotype-environment correlation have been described (Plomin et al., 1977). Passive gene-environment correlation refers to associations between child genotype and environmental exposure, that is, parents pass on genetic material as well as creating a home environment for their children based on their own (heritable) characteristics. Evocative genotype-environment correlation involves the child’s elicitation of environmental experience as a function of their genotype. Finally, active genotype-environment correlation suggests that children are active agents in their own socialization, including parenting, and this active manipulation of environmental experience is, at least in part, genetically determined. Despite a plethora of previous parenting research using behavioral genetic approaches to disentangle these effects (see Knafo & Jaffee, 2013), the field has thus far largely confounded parenting dimensions and their valence (positive or negative), potentially masking important underlying architectural distinctions. For the first time, we systematically examine positive and negative aspects of two parenting dimensions, parental feelings and parental control, in late childhood to early adolescence.

*Genetics of parenting*

Behavioral genetic studies of parenting are informative since they offer unique insights into the contributions of parent and child characteristics; the twin design can be useful in this context, comparing identical, or monozygotic (MZ), and fraternal, or dizygotic (DZ) twins. So-called parent-based twin designs compare the parenting behaviors of twin parents, with the premise that to the extent that parenting behaviors shown by MZ twin parents towards their offspring are more similar than those shown by DZ twin parents, genetically influenced parental characteristics are implicated. A child-based study, like the current one, has twin children as the focus for analyses. Here, the extent to which MZ twins are similarly parented compared to DZ twins is evidence of significant genetic influence on aspects of parenting, but suggests that parental behavior reflects genetically influenced characteristics of the child, such as their temperament.

*Parental Feelings and Control*

Subsequent theoretical and empirical perspectives on parenting have remained largely founded in Baumrind’s earlier work on parenting styles, which at its core, focused attention on two key aspects of parenting – responsiveness/warmth and demandingness/control (Baumrind, 1973). While researchers have distinguished aspects of parenting further, most notably in the area of parental control (e.g., Barber & Harmon, 2002), and have varied in their construct labels, these two broad dimensions have been endorsed through numerous studies that have sought to characterize them (e.g., Barber, Stolz, & Olsen, 2005; Cummings, Davies, & Campbell, 2000; Dodge, Petit & Bates, 1994; Maccoby & Martin, 1983; Rothbaum & Weisz, 1994). Here, we conceptualize these parenting dimensions as parental feelings (warmth, closeness, hostility, frustration) and parental control (discipline strategies such as remaining firm, and the use of physical discipline); these dimensions have shown robust modest to moderate associations to children’s outcomes (e.g., Parke & Buriel, 2006).

Reviews of behavioral genetic studies have noted that control aspects of parenting tend to yield low estimates of heritability while parental feelings yield moderate estimates (Kendler & Baker, 2007; Plomin, 1994; Rowe, 1981, 1983). To be clear, in child-based studies, these findings suggest that genetically influenced child characteristics may be more important for eliciting parental feelings than control. However, research has seldom distinguished between positive and negative parental feelings, and particularly between positive and negative control strategies. Blurring the positive and negative sides of feelings and control may mask important underlying foundations of parenting. Harsh discipline and effective supervision, for example, may not be opposite ends of a single continuum, and neither the same for hostility and warmth (see Pettit & Arsiwalla, 2008). Thus, we hypothesize that the underlying genetic architecture of these aspects of parenting may also be distinct. Specifically, following existing relevant family research (e.g., Rasbash, Jenkins, O’Connor, Tackett & Reiss, 2001; Kendler & Baker, 2007) as well as work outside the field (e.g., Baker et al., 1992; Baumeister, Bratslavsky, Finkenauer, & Vohs, 2001), we predict that negativity will show greater heritability than positivity across parental feelings and control as well as within parental feelings and within control.

For the first time in a large-scale, longitudinal, child-based twin study, we explicitly compare the genetic and environmental etiology of positivity and negativity, positive and negative feelings and positive and negative control.

Method

*Sample and Procedure*

The sampling frame for the current study was the Twins Early Development Study (TEDS), a population-based, longitudinal study of twins born in England and Wales in 1994-1996, recruited from UK birth records. TEDS has been demonstrated to be reasonably representative of the UK population (e.g., Harlaar, Spinath, Dale, & Plomin, 2005), acknowledging that literacy is a requirement for completing the questionnaires. For example, for families with 12 year data, 93.5% of the mothers in the sample self-reported ethnicity as ‘white’ and 43.6% of had qualifications at A’level or higher (the national educational qualification taken at 18 years in the UK); the equivalent UK population percentages for this generation are 93% white and 32% for A-levels or higher (Walker, Maher, Coulthard, Goddard, & Thomas, 2001). TEDS is described in detail elsewhere (see Haworth, Davis, & Plomin, 2013; Oliver & Plomin, 2007; Trouton, Spinath, & Plomin, 2002). Data at all ages were collected through parent questionnaires sent to families by mail. For the current study, at each age, we selected same-sex twin pairs only, and excluded all twin pairs where either twin had parent-reported medical or neurological conditions. Note that, due to differences in study procedures, data available differed at each age. Since this brief report does not include longitudinal analyses, we took the decision to include all appropriate data available at each age such that, while some families are included at all ages, some families are included at only one or two ages. Thus, the current study included 2260 twin pairs at age 9 (1202 MZ, 1058 DZ; 1034 boys, 1226 girls), 3850 twin pairs at age 12 (2027 MZ, 1823 DZ; 1752 boys, 2098 girls) and 2293 twin pairs at age 14 (1231 MZ, 1062 DZ; 1028 boys, 1265 girls). Zygosity was determined using parent ratings of physical similarity shown to be over 95% accurate when compared with DNA testing (see Price, Freeman, Craig, Petrill, Ebersole, & Plomin, 2000). For cases where zygosity was unclear, DNA testing was conducted to confirm zygosity. TEDS research was approved by the Institute of Psychiatry Ethics Committee, and all participants gave informed consent.

*Measures*

We generated eight scales from identical parent-report measures at child ages 9, 12, and 14 years of parental feelings and control. For feelings, we used an adapted, short form (seven items) of the Parental Feelings Questionnaire (PFQ; Deater-Deckard, 2000) and for control, a short (four item) discipline (parenting strategies) questionnaire adapted from Deater-Deckard et al. (1998). Item responses were rarely/never (0), sometimes (1), and often (2) and coding was reversed for items as necessary. Two standard composite measures were created at each age from the PFQ and Discipline questionnaires: Feelings from seven PFQ items, including the three positive (e.g., I feel close to my child) and four negative items (e.g., I feel frustrated by my child) and Control comprising four discipline items including two positive (e.g., I am firm and calm with him/her) and two negative (e.g., I tell him/her off or shout at him/her) items. In addition, scales were generated at each age to depict Negative Feelings (four negative PFQ items), Positive Feelings (three positive PFQ items), Negative Control (two negative discipline items) and Positive Control (two positive discipline items). We then created composite measures for all negative items (Negativity: four negative PFQ items, two negative discipline items) and for all positive items (Positivity: three positive PFQ items, two positive discipline items). For each scale, items were summed and averaged by the number of items with positive/negative valence and control/feelings elements as appropriate.

Although face validity for our scales is reasonable and appropriate for the hypothesis-driven nature of the current report, variable internal consistency for these scales was found, with reliabilities lower for scales with fewer items, as is to be expected (9 year: Feelings α = .68 (Negative feelings α = .75, Positive feelings α = .45), Control α = .44 (Negative control α = .33, Positive control α = .61), Negativity α = .74, Positivity α = .51; 12 year: Feelings α = .70 (Negative feelings α = .75, Positive feelings α = .50), Control α = .41 (Negative control α = .29, Positive control α = .63), Negativity α = .74, Positivity α = .52; 14 year: Feelings α = .72 (Negative feelings α = .75, Positive feelings α = .57), Control α = .36 (Negative control α = .25, Positive control α = .63), Negativity α = .75, Positivity α = .56).

*Analyses*

Data Preparation: The effects of children’s age and sex on measures were minimal, accounting for no more than 1% of the variance for any scale at any age. Nevertheless, standardized residual scores (controlling for age and sex) were used in the main analyses as is standard practice in twin studies to ensure that twin correlations are not artificially inflated due to the children being the same age and sex (McGue & Bouchard, 1984). All of the measures were somewhat skewed, with positive feelings the most skewed at each age (9 year: skew = 1.93; 12 year skew = 2.03; 14 year skew = 1.85; further details of skew on all measures are available on request). To avoid violation of the normality assumption implicit in the analyses involved, we applied Van der Waerden rank transformation to our measures (Lehmann, 1975). All analyses were conducted on these transformed data.

Twin Analyses: The classic twin method allows the decomposition of phenotypic variance into additive genetic (*A*), shared (*C*), and nonshared (*E*) environmental components (e.g., Plomin, DeFries, Knopik & Neiderhiser, 2013). Twin intraclass correlations (ICCs) provided an initial approximation of these ACE contributions. Heritability can be estimated as twice the difference between MZ and DZ ICCs, and shared environment as the difference between the MZ correlation and the heritability estimate. Estimates of the non-shared environmental component includes measurement error; it is the only source of variance responsible for MZ twin differences, and is thus estimated as the extent to which the MZ ICC is less than 1. Structural equation models provide more elegant estimates of these variance components as well as confidence intervals and model fit indices.

---FIGURE 1---

As well as fitting the standard univariate twin model (see Figure 1) for each of our scales across all ages, because the reliability of some scales was not optimal we further tested our hypotheses using the Common Pathway model.

---FIGURE 2---

In this model (see Figure 2), all common genetic and environmental variance is mediated through a single latent factor with age-specific factor loadings that allow influences on change over time ([Martin & Eaves, 1997](#_ENREF_5)). The common latent factor includes the common variance across the ages and is thus more reliable than the scales at each age. The common pathway model allows each phenotypic variable to be also influenced by specific environmental or genetic factors that are not shared across age, which we refer to as age-specific effects.

Model fitting and all subsequent analyses were done using OpenMx ([Boker et al., 2011](#_ENREF_2)), which uses minus twice log likelihood (-2LL) as the evaluation of the fit. To verify how well an alternative model fits the data a relative comparison of the fit statistic (-2LL) is made between the saturated model (the baseline model) and the model of interest (here, the common pathway model). Given that the difference between these statistics follows the chi-square distribution, we use the chi-square test with degrees of freedom equal to the difference between the number of parameters of each model. However, due to the sensitivity of this index to sample size, indices such as the Bayesian Information Criterion (BIC) ([Schwarz, 1978](#_ENREF_6)) or the Akaike Information Criterion (AIC) ([Akaike, 1987](#_ENREF_1)) are commonly used in behavioural genetic research ([e.g., Kendler, Gardner, & Lichtenstein, 2008](#_ENREF_3)), with suggestions that BIC performs better with larger sample sizes and more complex models ([Markon & Krueger, 2004](#_ENREF_4)). The lower the value of these indices, the better the balance of explanatory power and parsimony, indicates better fit.

Results

*Preliminary Analysis*

For information, descriptive statistics for all raw measures are included in Table 1, by sex and zygosity for each age.

---TABLE 1---

*Genetic Analysis*

As a first step in estimating genetic and environmental influence, twin intraclass correlations were calculated separately for the MZ and DZ twins (see Table 2). In all cases MZ twin similarity exceeds DZ twin similarity, indicating some genetic influence. The DZ twin correlations are greater than half the MZ twin correlations, indicating substantial shared environmental influence. Finally, the high MZ correlations indicate a minor role for nonshared environmental factors.

---TABLE 2---

Results from the model-fitting analyses (shown in Table 3) were highly similar to those gleaned from the intraclass correlations: genetic influence was significant for all scales, shared environment was substantial, and nonshared environment was negligible especially after discounting error of measurement. The focus of the present analysis is on heritability comparisons between feelings and control on the one hand and negativity and positivity on the other. As a prelude, the standard scales for PFQ Feelings and Control yielded a result to be expected from the literature: At all three ages, parental feelings were significantly more heritable than parental control, as indicated by the non-overlapping 95% confidence intervals for the Feelings and Control scales. The average heritability across the three ages was 42% for parental feelings and 13% for parental control. As an aside, such comparisons were robust for negative feelings (44%) compared to negative control (27%), and for positive feelings (26%) compared to positive control (6%).

Nonetheless, returning to our study focus, we found that all scales based on negative items are significantly more heritable than those based on positive items, again at all three ages. For example, average heritabilities across age for negative and positive feelings were 44% and 26%, respectively. Similarly, within the control scale, a scale based on the negative items was significantly more heritable than a scale based on the positive items at all ages, with average heritabilities across the three ages of 27% and 6%, respectively. Finally, creating scales for all the negativity items and all the positivity items regardless of whether they were on the feeling or control scale yielded significantly greater heritability for the negative scale than for the positive scale, with average heritabilities across the three ages of 44% and 12%, respectively.

---TABLE 3---

Acknowledging the fact that the reliabilities of some scales were less than desirable, we examined our hypothesis further by comparing saturated and common pathway models for each of our eight scales, feelings, control, negative feelings, positive feelings, negative control, positive control, negativity and positivity across the three ages, thus exploiting the stability of a latent factor for these analyses. The results for this common pathway model (see Figure 2) are shown in Table 4, which includes model fit statistics and parameter estimates. In every case, the common pathway model was found to fit the data at least equally as well as the saturated model, providing a statistical rationale for using the model.

---TABLE 4---

The pattern of heritability differences gleaned from the latent factors in these models remained the same as those for our univariate analyses of each scale at each age, adding further support for our hypothesis. That is, while latent factors for feelings and control differed significantly in their heritability in line with previous research (.38 and .08 respectively), there was additional emphasis on the distinctions between negativity and positivity. Specifically, for both feelings and control, negativity consistently yielded significantly higher heritability estimates than did positivity, a finding that held for the overall negativity and positivity latent factors (*h2* = .42 and.10 respectively). The pattern of heritability estimates for each of the individual scales calculated from the common pathway model did not differ from those reported in our main univariate analyses. Factor loadings suggested that, although age 12 had the strongest influence on the latent factors, all three ages were of similar magnitude. Note that, since they are not central to the focus of study, and in the interests of space in this brief report, the coefficients of the specific pathway are not reported here, but are available from the first author.

Discussion

For the first time, the current study explicitly examined the etiology of positive compared to negative aspects of parental feelings and control. A clear pattern of results emerged: the negative side of parenting showed significantly greater genetic influence than the positive side, regardless of whether parental feelings or control were assessed. Importantly, these results replicated across 9, 12 and 14 years, as well as using common pathway models across the three years in acknowledgment of the variable reliability of our measures and to add confidence in our findings, with all contrasts between positive and negative aspects of parenting significant in our sample.

Previous reviews have highlighted the lower heritability of parental control compared to parental feelings, suggesting that genetically influenced child traits may be more reflected in parental feelings than in parenting control strategies (Kendler & Baker, 2007; Plomin, 1994; Plomin & Bergeman, 1991). Further, it has been suggested that parental control encompasses more learned, socially-influenced parenting behaviors, whereas parental feelings are more strongly influenced by child characteristics such as temperament and behavior problems (e.g., Kendler & Baker, 2007). We too find that feelings show greater heritability than control over all. However, in line with our predictions, we have shown negative control to have greater genetic influence than positive control and negative feelings to have greater genetic influence than positive feelings. We argue that the ‘dark’ side of genetically driven child characteristics play a bigger role in eliciting parental negativity than do other child characteristics in eliciting positivity across feelings and control. For example, parental negativity encompassing hostility and harsh parenting may be more responsive to genetically driven challenging child temperament than positive features such as warmth and calmness are to less challenging traits. Although positivity versus negativity was not an explicit research question in previous studies, they lend support to our findings (Kendler & Baker, 2007). Further, a recent study using genetically informed social relations models indicates consistency in negative responses elicited by individuals in the family, and the importance of genetics for individual effects for negativity (Rasbash et al., 2011). Distinctions of parenting valence seem to be important for understanding family processes.

One of the earliest studies with relevant data was a study of 850 pairs of high-school twins (Loehlin & Nichols, 1976). In this study, parental control comprised items such as ‘stricter discipline’ which showed little genetic influence, yet there was one important exception - an item involving spanking, which showed considerable genetic influence. Indeed, other studies have similarly shown harsh parenting to reflect genetic influence (e.g., Jaffe, Caspi, Moffitt, Polo-Tomas, Price, & Taylor, 2004). This anomaly in heritability estimates could be due to harsh parenting involving more negative affect than other kinds of control (see Plomin & Bergeman, 1991). Notably, there is evidence outside the field of parenting of higher heritability for negative compared to positive affect (e.g., Baker et al., 1992). The reason why other studies have shown control scales to show less genetic influence than feelings may be that the control scales concerned have fewer negative-‘affective’ items. In the current study, our control scale includes negative affective aspects of control.

One caveat is critical here. In categorizing measures of parenting into positive versus negative valence, we do not include maltreatment. That is, the pattern we report includes aspects of harsh discipline such as yelling and spanking, but not abusive forms of parenting. One study that explicitly looked at this distinction found that while harsh discipline was moderately genetically influenced (25%), physical maltreatment was not (7%) (Jaffee et al., 2004). These findings suggest that children’s genetic influences are largely irrelevant for their vulnerability to maltreatment, and that characteristics of the perpetrator are what are important.

Although not a focus for this brief report, it is pertinent to highlight an additional finding of interest in the current study – that of a considerable shared environmental component both for control overall compared to feelings and for positive compared to negative aspects of parenting. In child-based designs like this one, shared environmental components indicate the extent to which parenting is consistent across children within the family. Our findings suggest that parents report being consistent in their discipline – particularly positive strategies – across members of the twin pair, as well in their feelings of warmth towards their children, once genetic similarities are accounted for.

We acknowledge limitations of the current study, such as the fact that the same parent rated both twins, likely to contribute to the high shared environment estimates in particular. The shared environmental estimates we find here mirror the genetic results in that higher shared environmental influence was found for control than feelings and for positivity than negativity. Although this may reflect genuine similarity in parental treatment strategies and warmth, it is possible that it is a parent reporting bias: parents may be more reluctant to say that they use different control strategies for their children or to admit to having more positive feelings about one twin than another.

We choose to use the simplicity of cross-sectional analyses to illustrate a novel path for parenting researchers within this brief report, with a view to offering a potential steer for more complex research. The robustness of our cross-sectional findings across age supports our perspective, but they are, by nature, limited. We acknowledge that the measures we use here are brief, and not designed for the question in hand; indeed the reliability of some our scales is variable, and this could have an impact on findings. Potentially, applying a more exploratory, factor analytical approach to our parenting variables would suggest they fall together differently. However, to evaluate our hypothesis driven, conceptual question in this large sample, it was necessary to generate our scales *a priori*. Further, our common pathways approach capitalized on the increased reliability of the latent factor, and yielded findings that additionally endorsed our hypotheses; moreover, the similarity of findings from our univariate and common pathways models suggest that the results from the former were robust to the lower reliabilities. However, we emphasize that more detailed measures across multiple methods and informants, as well as longitudinal study will be critical for teasing out the issues we have raised. While stressing this need for replication, we posit that our results highlight an important new angle for understanding the mutuality of child and parent influences on the parent-child relationship. In turn, our findings suggest several key directions for research, such as extending current multivariate behavioral genetic research that has examined child characteristics important for parenting (Knafo & Jaffee, 2013) to identify specific associations between positive and negative child characteristics with parental negativity and positivity across control and feelings. Using parent-based designs would also be of interest, to examine whether the ‘dark’ side of genetically influenced parent characteristics play similarly into the heritability of negativity and positivity in parent-child relationships. Further, another important question is the extent to which our findings tie in to contemporary intervention research that suggests that negative child characteristics may have a role in the effectiveness of interventions that aim to promote parental positivity (see Mash & Barkley, 2006; Sandler, Schoenfelder, Wolchik, & MacKinnon, 2011). Ultimately, studies that succeed in teasing out such child and parent effects could have key implications for informing parenting interventions. Finally, we suggest that our findings have considerable implications for work in other areas such as sibling and marital relationships, since negativity in these relations are likely to be similarly more genetically influenced than positivity, and thus related to genetically influenced characteristics of the members of the dyad.

We hope that the findings of this study will be a springboard for discussion, offering potential new perspectives on classic questions in developmental psychology. We argue that child characteristics may be especially important for influencing negative aspects of parenting, for control as well as for feelings, but emphasize that further work with multiple measures, methods and informants is needed. Better understanding how the environment is shaped by genetically driven individual differences in children’s characteristics is critical for basic science in developmental psychology, and, ultimately, for targeting interventions.

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Table 1: Descriptive statistics for all measures by zygosity.

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | *M (SD)* | | | | | | | | | | | | | | |
|  | 9 year | | | | | 12 year | | | | | 14 year | | | | |
|  | All | Boys | Girls | MZ | DZ | All | Boys | Girls | MZ | DZ | All | Boys | Girls | MZ | DZ |
| Measure | N=2257-2260 | n=1032-1034 | n=1224-1226 | n=1200-1203 | n=1055-1058 | N=3831-3850 | n=1743-1752 | n=2094-2099 | n=2016-2027 | n=1820-1825 | N=2286-2293 | n=1026-1028 | n=1258-1265 | n=1229-1231 | n=1056-1062 |
| Feelings | .87  (.55) | .91 (.56) | .84 (.55) | .88 (.56) | .86 (.55) | .82 (.58) | .87 (.59) | .79 (.57) | .81 (.58) | .84 (.57) | .82 (.61) | .85 (.61) | .80 (.60) | .82 (.60) | .82 (.61) |
| Control | 1.21 (.61) | 1.24 (.63) | 1.18 (.59) | 1.22 (.60) | 1.19 (.62) | 1.08 (.57) | 1.12 (.57) | 1.05 (.56) | 1.10 (.58) | 1.06 (.56) | 1.02 (.56) | 1.04 (.56) | 1.00 (.55) | 1.04 (.56) | .99 (.55) |
| Negative Feelings | .72 (.44) | .75 (.45) | .70 (.43) | .73 (.44) | .71 (.43) | .65 (.44) | .69 (.44) | .62 (.43) | .65 (.44) | .66 (.43) | .61 (.44) | .64 (.44) | .60 (.45) | .61 (.44) | .62 (.45) |
| Positive feelings | 1.85 (.26) | 1.84 (.27) | 1.86 (.25) | 1.85 (.26) | 1.85 (.26) | 1.83 (.29) | 1.82 (.30) | 1.84 (.28) | 1.84 (.8) | 1.82 (.29) | 1.79 (.32) | 1.79 (.33) | 1.80 (.31) | 1.80 (.32) | 1.79 (.32) |
| Negative control | .78 (.38) | .81 (.39) | .76 (.37) | .80 (.38) | .77 (.38) | .66 (.33) | .69 (.34) | .63 (.32) | .67 (.34) | .64 (.33) | .63 (.33) | .67 (.34) | .60 (.32) | .65 (.34) | .61 (.32) |
| Positive control | 1.57 (.44) | 1.56 (.44) | 1.58 (.43) | 1.57 (.43) | 1.57 (.44) | 1.57 (.44) | 1.57 (.44) | 1.58 (.44) | 1.57 (.45) | 1.58 (.43) | 1.61 (.44) | 1.63 (.44) | 1.60 (.44) | 1.61 (.43) | 1.62 (.44) |
| Negativity | 1.50 (.68) | 1.55 (.70) | 1.46 (.66) | 1.53 (.69) | 1.47 (.66) | 1.31 (.65) | 1.38 (.66) | 1.25 (.63) | 1.32 (.65) | 1.30 (.64) | 1.25 (.67) | 1.31 (.68) | 1.20 (.65) | 1.26 (.67) | 1.23 (.66) |
| Positivity | 3.42 (.55) | 3.40 (.57) | 3.44 (.53) | 3.43 (.54) | 3.42 (.56) | 3.40 (.56) | 3.39 (.56) | 3.41 (.56) | 3.41 (.57) | 3.40 (.56) | 3.41 (.59) | 3.41 (.59) | 3.40 (.58) | 3.40 (.58) | 3.41 (.59) |

*Note:* *M*=mean; *SD*=standard deviation; MZ=monozygotic; DZ=dizygotic; sample sizes given are twin pairs; for all measures, scores represent the average item score such that a mean of 0 indicates an average item score representing ‘rarely/never’ and a mean of 2 indicates an average item score representing ‘often’.

Table 2: Monozygotic (MZ) and dizygotic (DZ) intraclass correlations (ICCs) and model-fitting estimates of genetic (A), shared (C) and nonshared (E) environmental variance components across scales and ages

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | 9 year | | | | | 12 year | | | | | 14 year | | | | |
|  | *ICC* | |  |  |  | *ICC* | |  |  |  | *ICC* | |  |  |  |
| Measure | MZ | DZ | A  (CI) | C  (CI) | E  (CI) | MZ | DZ | A  (CI) | C  (CI) | E  (CI) | MZ | DZ | A  (CI) | C  (CI) | E  (CI) |
| Feelings | 0.91 | 0.68 | .47  (.41-.53) | .45  (.38-.50) | .09  (.08-.10) | 0.90 | 0.72 | .35  (.31-.39) | .55  (.51-.59) | .10  (.10-.11) | 0.90 | 0.69 | .43  (.37-.49) | .47  (.41-.53) | .10  (.09-.11) |
| Control | 0.97 | 0.91 | .13  (.11-.15) | .85  (.83-.86) | .03  (.02-.03) | 0.96 | 0.91 | .12  (.11-.14) | .84  (.83-.86) | .04  (.03-.04) | 0.97 | 0.90 | .14  (.12-.16) | .83  (.81-.85) | .03  (.03-.04) |
| Negative Feelings | 0.91 | 0.66 | .50  (.44-.57) | .41  (.34-.47) | .09  (.08-.10) | 0.89 | 0.69 | .39  (.35-.44) | .49  (.45-.53) | .12  (.11-.12) | 0.9 | 0.67 | .44  (.39-.51) | .45  (.39-.51) | .10  (.09-.11) |
| Positive feelings | 0.91 | 0.79 | .25  (.21-.29) | .67  (.62-.70) | .09  (.08-.09) | 0.91 | 0.79 | .24  (.21-.27) | .67  (.64-.70) | .09  (.08-.10) | 0.92 | 0.77 | .30  (.26-.35) | .62  (.57-.66) | .08  (.07-.09) |
| Negative control | 0.96 | 0.81 | .31  (.28-.35) | .65  (.61-.68) | .04  (.04-.05) | 0.94 | 0.81 | .26  (.24-.29) | .68  (.65-.70) | .06  (.06-.07) | 0.94 | 0.82 | .24  (.21-.28) | .70  (.66-.73) | .06  (.05-.06) |
| Positive control | 0.99 | 0.97 | .04  (.04-.05) | .95  (.94-.95) | .01  (.01-.01) | 0.98 | 0.94 | .08  (.07-.09) | .90  (.89-.91) | .02  (.02-.02) | 0.99 | 0.96 | .06  (.05-.07) | .93  (.92-.94) | .01  (.01-.01) |
| Negativity | 0.93 | 0.69 | .47  (.42-.53) | .46  (.40-.51) | .07  (.06-.07) | 0.91 | 0.71 | .41  (.37-.45) | .51  (.46-.55) | .09  (.08-.10) | 0.92 | 0.7 | .43  (.38-.49) | .49  (.43-.54) | .08  (.07-.09) |
| Positivity | 0.96 | 0.92 | .08  (.07-.10) | .88  (.86-.89) | .04  (.03-.04) | 0.96 | 0.9 | .10  (.09-.12) | .85  (.84-.87) | .05  (.04-.05) | 0.97 | 0.89 | .16  (.14-.19) | .80  (.78-.82) | .03  (.03-.04) |

*Note:* *CI*=95% confidence intervals

Table 3: Model fit statistics for all univariate models using OpenMx ([Boker et al., 2011](#_ENREF_2))

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | 9 year | | | | 12 year | | | | 14 year | | | |
| Measure | -2LL | df | AIC | BIC | -2LL | df | AIC | BIC | -2LL | df | AIC | BIC |
| Feelings | 9979.73 | 4514 | 951.73 | -29571.33 | 17026.28 | 7682 | 1662.28 | -50282.35 | 10129.32 | 4568 | 993.32 | -29894.87 |
| Control | 7379.56 | 4508 | -1636.44 | -32118.92 | 13087.14 | 7685 | -2282.86 | -54247.78 | 7879.48 | 4579 | -1278.52 | -32241.09 |
| Negative Feelings | 10030.12 | 4514 | 1002.12 | -29520.93 | 17346.69 | 7684 | 1978.69 | -49979.47 | 10262.71 | 4574 | 1114.71 | -29814.05 |
| Positive feelings | 9438.07 | 4517 | 404.07 | -30139.26 | 16232.04 | 7694 | 844.04 | -51181.74 | 9703.88 | 4571 | 561.88 | -30346.59 |
| Negative control | 8705.75 | 4512 | -318.25 | -30827.78 | 15266.72 | 7689 | -111.28 | -52103.25 | 9125.64 | 4583 | -40.36 | -31029.98 |
| Positive control | 5529.89 | 4514 | -3498.14 | -34021.19 | 11253.52 | 7687 | -4120.48 | -56098.92 | 5538.73 | 4579 | -3619.27 | -34581.84 |
| Negativity | 9602.69 | 4509 | 584.69 | -29904.55 | 16700.89 | 7668 | 1364.89 | -50485.08 | 9923.17 | 4574 | 775.17 | -30153.59 |
| Positivity | 7695.63 | 4513 | -1330.37 | -31846.66 | 13768.72 | 7676 | -1583.28 | 53487.34 | 8070.73 | 4567 | -1063.27 | -31944.70 |

Table 4: Model fit statistics and parameter estimates for common pathway models using OpenMx ([Boker et al., 2011](#_ENREF_2))

|  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | Saturated model | | | | Common Pathway | | | | Coefficients | | |
| Latent Factor | -2LL | df | AIC | BIC | -2LL | df | AIC | BIC | A  (CI) | C  (CI) | E  (CI) |
| Feelings | 34764.07 | 16722 | 1320.07 | -111751.80 | 34817.73 | 16759 | 1299.73 | -112022.34 | .38  (.33-.44) | .58  (.52-.63) | .04  (.03-.05) |
| Control | 27022.39 | 16730 | -6437.61 | -119563.58 | 27089.87 | 16767 | -6444.13 | -119820.29 | .08  (.06-.10) | .91  (.89-.93) | .01  (.00-.01) |
| Negative Feelings | 34973.88 | 16730 | 1513.89 | -111612.09 | 35020.71 | 16767 | 1486.71 | -111889.45 | .44  (.39-.51) | .51  (.45-.57) | .04  (.04-.05) |
| Positive feelings | 34240.45 | 16740 | 760.45 | -112433.14 | 34272.38 | 16777 | 718.38 | -112725.40 | .29  (.23-.35) | .67  (.62-.72) | .04  (.03-.05) |
| Negative control | 31429.89 | 16742 | -2054.11 | -115261.23 | 31482.87 | 16779 | -2075.13 | -115532.44 | .21  (.17-.24) | .78  (.75-.82) | .01  (.01-.01) |
| Positive control | 21833.77 | 16738 | -11642.23 | -124822.30 | 21893.82 | 16775 | -11656.18 | -125086.44 | .03  (.02-.05) | .96  (.95-.98) | .01  (.00-.01) |
| Negativity | 33415.55 | 16709 | -2.45 | -112986.43 | 33471.51 | 16746 | -20.49 | -113254.65 | .42  (.37-.47) | .54  (.49-.59) | .04  (.03-.04) |
| Positivity | 28688.11 | 16714 | -4739.89 | -117757.67 | 28730.80 | 16751 | -4771.20 | -118039.17 | .10  (.07-.13) | .89  (.86-.91) | .02  (.01-.02) |

Figure 1: Univariate individual differences model fitting: A, additive genetic influence; C, shared environment; E, non-shared environment; paths a, c and e, effects of A, C and E on the quantitative trait. Genetic relatedness, or the genetic correlation (ra) is 1.0 for MZ twins and 0.5 for DZ twins; environmental correlation (rc) is assumed to be 1.0 both for MZ and DZ twins.

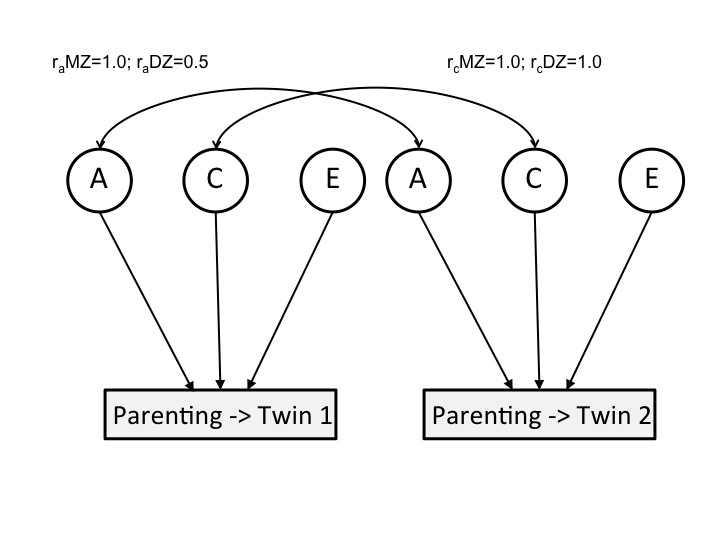


Figure 2: Common pathway model: al, cl, and el represent additive genetic, shared environment and non-shared environmental components of the latent factor variance across the three ages; paths f7, f12 and f14 are the factor loadings from measures at each age; as, cs, es are the specific additive genetic, shared environmental and non-shared environmental influences on scales at each age.

